

Introduction

An ideal simultaneous confidence intervals (SCI) for sparse linear model

- 1) should be as tight as possible and achieve the nominal confidence level simultaneously (coverage probability, the width of intervals of nonzero and zero coefficients);
- 2) **should be able to imply the variable selection results** in a way that the truly relevant and irrelevant coefficients could have nonzero and zero width intervals, respectively.

We propose a general approach to construct simultaneous confidence intervals based on entire solution paths and residual bootstraps.

General Approach for Constructing SCI

We define an **outlyingness score** for each bootstrap estimator to measure the relative location of a bootstrap estimator among all B bootstrap estimators as follow:

$$O^{(b)} = g(\hat{\beta}^{(b)}) = (o_1^{(b)}, \dots, o_d^{(b)}) \in \mathbb{R}^{+d}, b \in 1, \dots, B.$$

Then, we can rule out α percent of outlying bootstrap estimators among all to construct the simultaneous confidence intervals with confidence level $1-\alpha$.

Two special instances of outlyingness score:

$$1. O^{F,(b)} = (o^{F,(b)}) = g^F(\hat{\beta}^{(b)}) = \hat{F}(\gamma_b, \gamma_f) = \frac{(RSS_{\gamma_b} - RSS_{\gamma_f}) / (df_{\gamma_b} - df_{\gamma_f})}{RSS_{\gamma_f} / df_{\gamma_f}}$$

$$2. O^{\text{MaxMin},(b)} = (o_{\max}^{(b)}, o_{\min}^{(b)}) = g^{\text{MaxMin}}(\hat{\beta}^{(b)}) = \left(\max_{j \in \{1, \dots, p\}} \left(\frac{\hat{\beta}_j^{(b)} - \tilde{\beta}_j}{s.e. \hat{\beta}_j} \right), \min_{j \in \{1, \dots, p\}} \left(\frac{\hat{\beta}_j^{(b)} - \tilde{\beta}_j}{s.e. \hat{\beta}_j} \right) \right)$$

Procedure: Simultaneous Confidence Intervals

Step 1: Apply residual bootstrap for variable selection to obtain: $\{\hat{\beta}^{(b)}\}_{b=1}^B$;

Step 2: Construct outlyingness score: $O^{(b)} = (o_1, o_2, \dots, o_d) = g(\hat{\beta}^{(b)}) \in \mathbb{R}^{+d}$;

Step 3: Construct a set $\mathcal{A}_\alpha \subset \{1, \dots, B\}$:

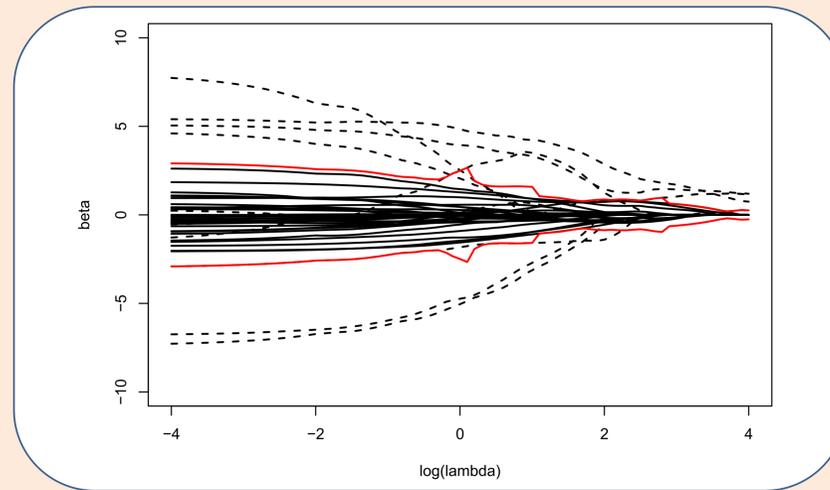
$$\mathcal{A}_\alpha = \{b \in \{1, \dots, B\}; o_i^{(b)} \leq q_i(1 - \frac{\alpha}{d}), i = 1, \dots, d\},$$

where $q_i(1 - \frac{\alpha}{d})$ is $(1 - \frac{\alpha}{d})$ quintile of o_i ;

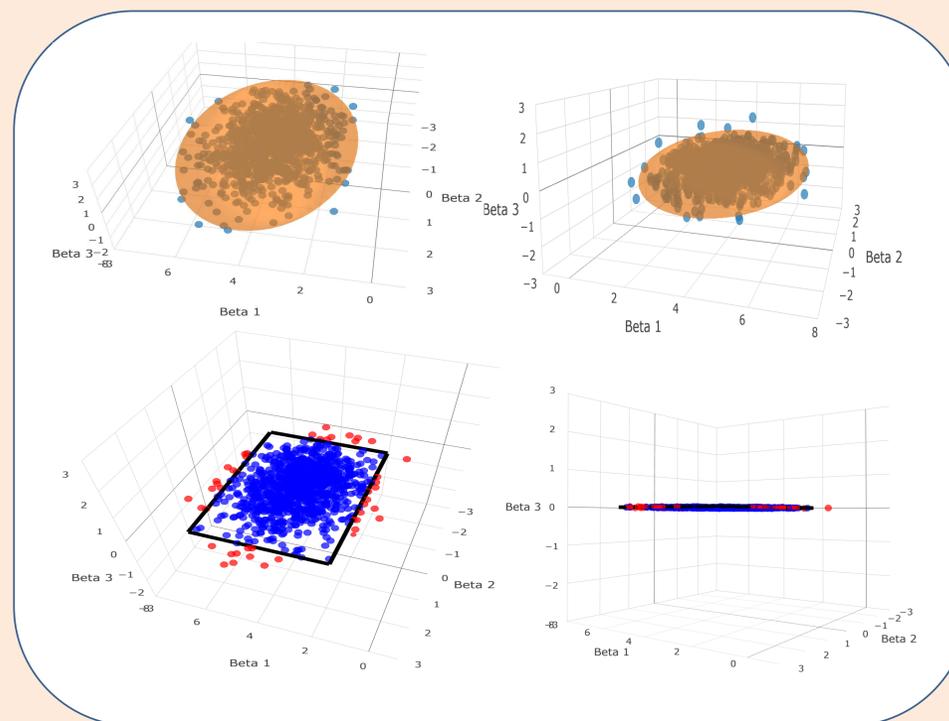
Step 4: Construct the simultaneous confidence intervals (SCI) as:

$$SCI(1 - \alpha) = \left\{ \beta \in \mathbb{R}^p; \min_{b \in \mathcal{A}_\alpha} \beta_j^{(b)} \leq \beta_j \leq \max_{b \in \mathcal{A}_\alpha} \beta_j^{(b)}, j = 1, \dots, p \right\}.$$

Selection by Partitioning the Solution Paths (SPSP)



Geometrical Differences: Proposed vs. Debiased



Theoretical Results

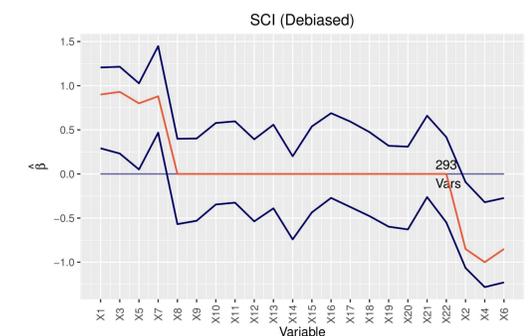
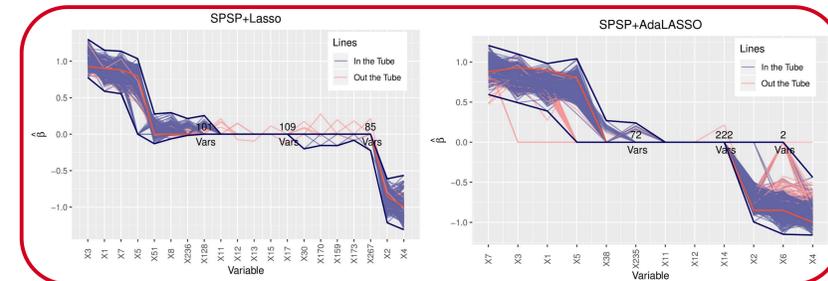
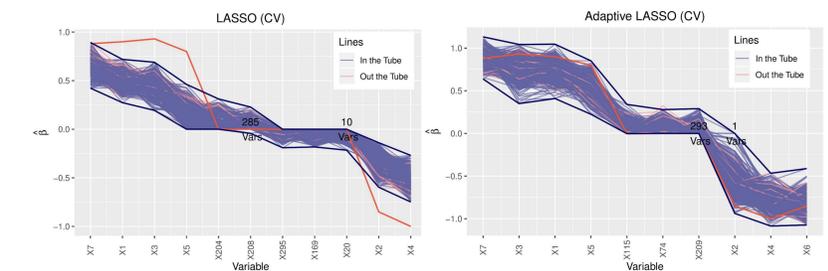
Theorem: Under mild assumptions, for $\alpha \in (0, 1)$ and all $\beta \in \mathbb{R}^p$, we have

$$\mathbf{P}(\beta \in SCI_{n,(1-\alpha)}) \rightarrow 1 - \alpha \text{ as } n \rightarrow \infty.$$

Simulation Example

- ❖ $p=300, n=200$
- ❖ $\beta^* = (0.9, -0.85, 0.93, -1, 0.8, -0.85, 0.88)$
- ❖ Remaining coefficients equal zero
- ❖ Correlation between X_i and X_j is $0.5^{|i-j|}$

SPSP vs. Cross-Validation vs. Debiased



SCI	W.Nzero	W.Zero	Cover Pr	Avg Card	Med Card	Std Card
SPSP+AdaLasso(MaxMin)	0.60	0.04	96.50	68.31	59.00	51.66
SPSP+AdaLasso(F)	0.61	0.06	98.50			
SPSP+Lasso(MaxMin)	0.92	0.19	96.50	734.19	770.50	150.75
SPSP+Lasso(F)	0.92	0.19	96.50			
AdaLasso(MaxMin)	0.64	0.21	66.00	949.24	950.00	1.56
AdaLasso(F)	0.64	0.21	65.50			
Lasso(MaxMin)	0.54	0.25	0.00	950.00	950.00	0.00
Lasso(F)	0.54	0.25	0.00			
True model(MaxMin)	0.45	0.00	92.50	1.00	1.00	0.00
True model(F)	0.46	0.00	99.50			
SCI(Debiased)	0.97	0.97	98.00			